

Confidence-Calibrated LLM Pipeline for Adverse Drug Reaction Detection from Clinical Instant Messaging: Development and Temporal Validation

Dongxu Wang^{1,†}, Zihong Lu^{1,†}, Wenbo Yuan¹, Kaiqiang Yuan²,
Di Yin¹, Ying Yao^{1,*}, Sunmin Jiang^{1,*}

¹Department of Pharmacy, Affiliated Women’s Hospital of Jiangnan University, Wuxi, China

²Guangzhou Pinyi Information Technology Co., Ltd., Guangzhou, China

[†]These authors contributed equally to this work.

*Corresponding authors: Ying Yao, Sunmin Jiang

Abstract

Objective: To develop and temporally validate a confidence-calibrated large language model (LLM) pipeline for adverse drug reaction (ADR) detection, entity extraction, and causality assessment from clinical instant messaging (IM).

Materials and Methods: Development proceeded in two phases. In the pilot phase, we evaluated a hybrid rule-LLM pipeline on 450 simulated messages (Fleiss’ $\kappa=0.719$) and 1,792 real messages from a hospital pharmacy WeChat group. In the validation phase, we constructed a 2,023-message gold standard (2021–2025) annotated by five pharmacists via blind review, with temporal split into development (<2024 , $n=1,277$) and locked test (≥ 2024 , $n=746$) sets. The pipeline uses Qwen 3.5 Plus with 1–10 confidence scoring. Evaluations included calibration analysis, multi-turn context ablation, cross-model robustness (four LLMs), a fine-tuned BERT baseline, specificity on 1,000 synthetic medical-but-non-ADR messages, and Naranjo/IMCT causality assessment on 200 cases.

19 **Results:** The pipeline improved from pilot F1=0.906 to temporally validated
20 F1=0.970 (P=0.944, R=0.997), matching the BERT baseline. Confidence scores
21 were well-calibrated (expected calibration error=0.030). Single-message classifica-
22 tion outperformed multi-turn approaches ($p < 0.001$). Entity extraction achieved
23 lenient F1 of 0.750 (drug) and 0.738 (symptom). All causality instruments showed
24 no agreement with pharmacist consensus (intraclass correlation coefficient ≤ -0.236).

25 **Discussion:** The pipeline achieves near-expert ADR detection with calibrated
26 uncertainty, but causality assessment is limited by the information content of brief
27 IM messages.

28 **Conclusion:** Confidence-calibrated LLM classification is effective for IM-based
29 ADR screening (projected positive predictive value=71% at 10% prevalence, based
30 on synthetic negative controls). Single-center findings require multi-site validation.
31 Causality assessment should be deferred to formal review.

32 **Keywords:** adverse drug reaction, pharmacovigilance, large language model, clinical
33 instant messaging, confidence calibration, temporal validation

34 **Lay Summary:** Clinical pharmacists discuss suspected drug side effects in hospital mes-
35 saging groups, but these valuable safety signals are not systematically captured. We devel-
36 oped an AI system that automatically identifies adverse drug reaction reports from these
37 conversations with an F1-score of 0.970, matching the performance of a purpose-trained
38 model without requiring task-specific training data. The system provides calibrated con-
39 fidence scores that enable priority-based review. While effective for detection, automated
40 causality assessment from brief messages remains infeasible, indicating that these systems
41 should serve as screening tools with human expert follow-up.

1 Introduction

Adverse drug reactions (ADRs) account for 5–8% of hospital admissions and are a leading cause of preventable patient harm [1–3]. Although spontaneous reporting systems remain central to post-marketing pharmacovigilance [4], up to 94% of ADRs go unreported, primarily because voluntary reporting requires clinicians to interrupt workflow and complete structured forms [5–7].

Clinical instant messaging (IM) platforms—particularly WeChat and WeCom in China—have become integral to hospital communication [8]. Pharmacists routinely discuss suspected adverse reactions through these channels, generating pharmacovigilance signals as a byproduct of clinical practice. Yet scoping reviews of natural language processing (NLP)-based adverse event detection from electronic health records [9, 10] and social media [11, 12] have identified no system utilizing IM-derived clinical text.

Large language models (LLMs) have demonstrated strong performance in adverse event extraction from clinical notes [13, 14], surveillance reports [15], and social media [16], though challenges persist in domain variability and hallucination [17]. No prior work has applied LLMs to clinical IM conversations or attempted automated causality assessment from informal conversational text.

In a pilot study, we developed a hybrid rule-LLM pipeline that achieved F1=0.906 on 450 simulated messages and F1=0.905 on 1,792 real clinical messages, with zero false positives. However, the pilot evaluation relied on a single annotator’s simulated benchmark, lacked temporal separation, and did not validate entity extraction or causality assessment. The present study extends this work through a comprehensive two-phase development and validation design with five contributions:

1. **Rigorous temporal validation** with a 2,023-message gold standard annotated by five pharmacists, strict development/test separation, and a fine-tuned BERT baseline.
2. **Confidence-calibrated classification with architectural ablation**, evaluating calibration quality and testing whether two-pass re-examination improves over single-pass classification.

- 70 **3. Multi-turn context analysis and entity extraction validation** against phar-
71 macist annotations across four independent Chinese LLMs.
- 72 **4. Specificity and deployment feasibility**, including prevalence-adjusted positive
73 predictive value for lower-prevalence settings.
- 74 **5. Systematic causality assessment** using three instrument variants, characterizing
75 the limits of IM-based causality scoring.

76 **2 Materials and Methods**

77 **2.1 Study Design Overview**

78 This study followed a two-phase design. Phase 1 (pilot) developed the pipeline architec-
79 ture and validated it on simulated and real-world messages. Phase 2 (temporal validation)
80 constructed a multi-annotator gold standard and conducted comprehensive evaluation
81 with confidence calibration, multi-turn analysis, cross-model robustness, and causality
82 assessment.

83 **2.2 Phase 1: Pilot Validation**

84 **2.2.1 Simulated Benchmark**

85 A clinical pharmacist with 5 years of experience authored 450 messages based on authen-
86 tic drug safety scenarios, stratified by difficulty: easy (n=110; explicit drug-symptom
87 co-mentions), medium (n=165; brand names, abbreviations, colloquial phrasing), and
88 hard (n=175; implicit causality, ambiguous referents, lab values). Five independent phar-
89 macists validated clinical realism (Fleiss' $\kappa=0.719$, mean realism rating 3.81/5) [18].

90 **2.2.2 Real-World Pilot Data**

91 A non-overlapping set of 1,792 authentic messages from the hospital's pharmacovigilance
92 WeChat group (March 2024–February 2025; 1,372 ADR+, 420 ADR-), annotated by
93 a single pharmacist. These messages are temporally and numerically distinct from the

94 Phase 2 dataset; none were included in Phase 2 development or test sets. A negative con-
95 trol corpus of 3,897 messages from a separate pharmacy quality-control group (containing
96 drug names in non-ADR contexts) served as the specificity test set.

97 **2.3 Phase 2: Temporal Validation Datasets**

98 **2.3.1 Data Source**

99 The Phase 2 dataset comprises 2,023 messages from a pharmacovigilance-dedicated WeChat
100 group at Wuxi Maternity and Child Health Care Hospital, a tertiary obstetrics and gyne-
101 cology hospital, collected between September 2021 and March 2025. Messages represent
102 the complete unfiltered archive from this group.

103 **2.3.2 Gold Standard Construction**

104 Five clinical pharmacists (each with ≥ 3 years of experience) independently annotated
105 each message as ADR-positive or ADR-negative using blind review. An ADR-positive
106 message was defined as containing: (1) an identifiable drug name; (2) a described adverse
107 symptom or clinical sign; and (3) an explicit or implied temporal association between
108 drug administration and symptom onset. A calibration session using 50 pilot messages
109 preceded independent annotation. Inter-annotator agreement was Fleiss' $\kappa=0.697$ (95%
110 CI: 0.677–0.717), indicating substantial agreement [19], with full 5/5 agreement on 69.1%
111 of messages (n=1,398), 4/5 agreement on 17.4% (n=351), and 3/5 agreement on 13.5%
112 (n=274).

113 We determined the final label through a two-stage process: initial majority vote
114 ($\geq 3/5$), followed by systematic adjudication by two senior pharmacists (not involved
115 in initial annotation) who reviewed all cases where the majority vote conflicted with the
116 objective annotation criteria. Adjudication was unidirectional (ADR- to ADR+ only),
117 correcting cases where messages contained all three required elements (drug name, adverse
118 symptom, and temporal association) yet were labeled negative by majority vote, primar-
119 ily involving dechallenge patterns, abbreviated drug names, and chemotherapy laboratory
120 abnormalities. In total, 734 labels were changed (439 in the development set, 295 in the

121 test set): 497 had 0/5 positive votes, 118 had 1/5, and 119 had 2/5. The high cor-
122 rection rate among 0/5 cases reflects messages where all annotators overlooked implicit
123 ADR patterns (e.g., dechallenge narratives without explicit drug names) that met the pre-
124 specified criteria upon expert review. Adjudication criteria were defined independently
125 of any pipeline output; all corrections with original votes and rationale are documented
126 in the code repository. The final dataset contains 1,692 ADR-positive (83.6%) and 331
127 ADR-negative (16.4%) messages, compared with 958 (47.4%) positive by initial majority
128 vote alone.

129 **2.3.3 Temporal Split**

130 We split messages by timestamp at January 2024 to prevent data leakage. We chose this
131 cutoff a priori based on two criteria: (1) achieving approximately 60/40 development/test
132 proportions, and (2) ensuring the test set captured the most recent clinical patterns,
133 including any temporal drift in reporting conventions or drug formulary changes:

- 134 • **Development set** (< January 2024): 1,277 messages (1,054 ADR+, 223 ADR-)
- 135 • **Test set** (\geq January 2024): 746 messages (638 ADR+, 108 ADR-)

136 All optimization experiments used only the development set; the test set was accessed
137 once for final validation.

138 **2.3.4 Missing Data**

139 No messages were excluded for missing or incomplete text content. All 2,023 messages in
140 the pharmacovigilance group archive contained extractable text and were included in the
141 annotation process.

142 **2.3.5 Conversation Structure**

143 We grouped messages into 401 multi-message conversations (≥ 2 messages): 245 (791
144 messages) in the development set and 156 (393 messages) in the test set, used for multi-
145 turn context experiments.

146 **2.3.6 Medical Negative Controls**

147 To evaluate specificity, we generated 1,000 synthetic messages across 11 categories of
148 medical-but-non-ADR content (e.g., confounding patterns where drug-symptom co-occurrence
149 reflects expected pharmacological effects, n=120; disease symptoms without drug involve-
150 ment, n=90; dosage consultations, n=90). Templates incorporated realistic drug names
151 and clinical parameters, achieving a 63.5% drug keyword rate and 4.1% drug-symptom
152 co-occurrence rate.

153 **2.4 Pipeline Architecture**

154 The pipeline uses a two-layer design developed in Phase 1: a deterministic rule-based pre-
155 filter and an LLM classifier. The sole input to all configurations is the raw text content of
156 each individual IM message; no structured metadata (sender identity, timestamp, patient
157 demographics) or external knowledge bases are used as predictors.

158 **2.4.1 Rule Layer**

159 The rule engine matches messages against vocabularies of 140+ drug names (brand,
160 generic, and abbreviated) and 60+ symptom patterns. Messages matching both are clas-
161 sified as ADR-positive; those matching neither are classified as ADR-negative; remaining
162 messages proceed to the LLM.

163 **2.4.2 LLM Layer with Confidence Scoring**

164 The LLM layer uses Qwen 3.5 Plus [20] via the Aliyun DashScope API. A few-shot
165 prompt provides classification criteria with domain-specific guidance (chemotherapy lab-
166 oratory abnormalities, abbreviated drug names), four annotated examples, and instruc-
167 tions to output JSON containing: binary classification, confidence score (1–10 integer),
168 drug name, symptoms, patient identifier, and reasoning. The prompt was optimized on
169 the development set and locked before test evaluation.

170 **2.4.3 Two-Pass Architecture**

171 We evaluated a confidence-gated two-pass design where intermediate-confidence cases
 172 (between thresholds θ_L and θ_H) undergo error-pattern-specific re-examination. Pass 2
 173 incorporates targeted guidance for chemotherapy lab values, short texts (<20 characters),
 174 and confounding patterns. Thresholds ($\theta_H=9$, $\theta_L=2$) were optimized on the development
 175 set.

176 **2.4.4 Pipeline Configurations**

177 We compared three variants: **rule-only** (no LLM), **LLM-only** (no rule pre-filter), and
 178 **hybrid** (rule pre-filter with LLM fallback). Figure 1 illustrates the pipeline architecture.

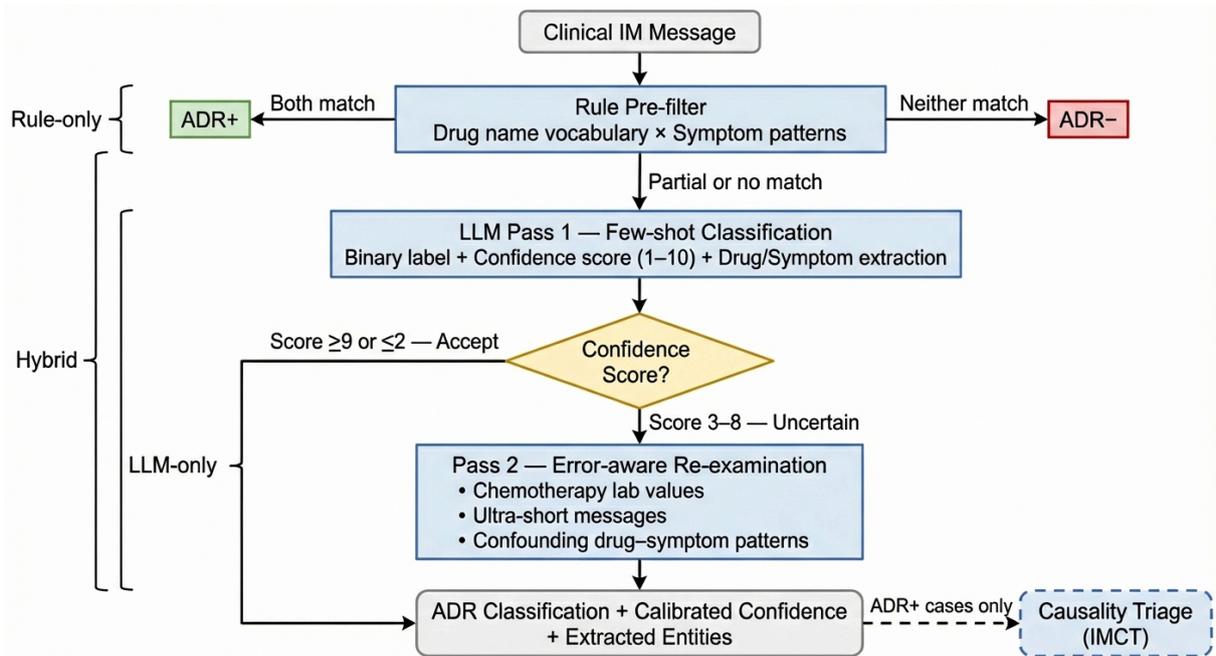


Figure 1: Pipeline architecture for ADR detection from clinical IM messages. Three configurations were evaluated: rule-only (deterministic pre-filter), LLM-only (few-shot classification with confidence scoring), and hybrid (rule pre-filter with LLM fallback). Uncertain cases from Pass 1 are routed through error-pattern-specific re-examination in Pass 2.

179 **2.5 Causality Assessment**

180 **2.5.1 Automated Naranjo Scale**

181 For ADR-positive cases, we implemented LLM-based Naranjo assessment [21] using the
 182 10-item scale (Definite ≥ 9 , Probable 5–8, Possible 1–4, Doubtful ≤ 0). The LLM receives

183 conversation context and outputs per-question answers with evidence and information
184 source. Three pharmacists independently scored 200 cases as gold standard, with consen-
185 sus by majority vote.

186 **2.5.2 IM Causality Triage**

187 Given the Naranjo scale’s structural mismatch with IM data (Section 3.9), we designed a
188 3-question IM Causality Triage (IMCT) focusing on IM-answerable dimensions:

- 189 **1. Q1: Known drug-ADR association?** Whether the drug-symptom pair is a
190 recognized adverse reaction.
- 191 **2. Q2: Temporal evidence?** Whether temporal markers indicate drug preceded
192 symptom.
- 193 **3. Q3: Dechallenge evidence?** Whether symptom improvement after discontinua-
194 tion is described.

195 Triage categories: HIGH (Q1=Yes AND [Q2 or Q3]=Yes), MEDIUM (partial evidence),
196 LOW (all Unknown). Gold standard mappings used existing Naranjo annotations (Q2→Naranjo Q2,
197 Q3→Naranjo Q3) with Q1 assessed against pharmacological knowledge.

198 **2.6 Entity Extraction Evaluation**

199 For ADR-positive messages with pharmacist annotations (n=958), extraction accuracy
200 was evaluated using strict matching (normalized exact match after alias resolution) and
201 lenient matching (token-overlap F1 with substring matching), stratified by split and an-
202 notator agreement level.

203 **2.7 Experimental Design**

204 Experiments were organized into four groups:

205 **Group A: Core validation.** Architecture comparison (Exp1), prompt optimization
206 (Exp2), and locked test set validation (Exp3).

207 **Group B: Multi-turn context.** Context window ablation on the development subset
208 (Exp4; windows 0, 1, 3, 5, full) and temporal validation on the test subset (Exp5).

209 **Group C: Causality.** Full Naranjo (Exp6), simplified 5-question Naranjo (Exp7), and
210 IMCT (Exp10) on 200 cases.

211 **Group D: Robustness and extraction.** Inference stability (Exp8; five runs, tem-
212 perature=0.1), error taxonomy (Exp9), medical negative control specificity, cross-model
213 comparison (Qwen 3.5 Plus, DeepSeek V3.2 [22], Kimi K2.5, GLM-5 [23]), supervised Bidi-
214 rectional Encoder Representations from Transformers (BERT)-base-Chinese baseline [24]
215 (five-fold stratified CV on the development set; max sequence length=128; batch size=16;
216 AdamW optimizer, lr= 2×10^{-5} , weight decay=0.01; early stopping with patience=3 on
217 validation loss; maximum 10 epochs), confidence calibration analysis, two-pass ablation,
218 and entity extraction evaluation.

219 **2.8 Statistical Analysis**

220 We assessed classification performance using precision, recall, F1, and specificity with
221 95% bias-corrected and accelerated (BCa) bootstrap confidence intervals (10,000 resam-
222 ples, seed=42) [25]. McNemar’s test with continuity correction was used for paired com-
223 parisons; exact p -values are reported alongside χ^2 statistics. Where multiple pairwise
224 comparisons were conducted within an experiment (e.g., context window ablation), Bon-
225 ferroni correction was applied and adjusted significance thresholds are noted. Calibration
226 was evaluated using expected calibration error (ECE, 10 bins) and Brier score. Causality
227 agreement used intraclass correlation coefficient [ICC(2,1)], Cohen’s κ , and mean absolute
228 error (MAE). Entity extraction used strict and lenient matching stratified by split and
229 agreement level. Inference stability used F1 coefficient of variation (CV) and unanimous
230 agreement rate. All analyses used Python 3.13.

2.9 Ethics

This study was a retrospective analysis of de-identified clinical instant messages with no patient interaction or intervention. All messages were de-identified prior to analysis: patient identifiers were replaced with hashed anonymous codes, patient names were substituted with uniform placeholder tokens, and sender identities were anonymised with sequential codes. The institutional review board of Wuxi Maternity and Child Health Care Hospital determined that formal ethics approval was not required. Clinical pharmacists participating in the annotation study provided verbal informed consent. For the retrospective IM data, individual consent was waived given the de-identified, retrospective nature of the analysis. LLM inference used the Aliyun DashScope API under data processing agreements compliant with China’s Personal Information Protection Law [26].

3 Results

3.1 Phase 1: Pilot Validation

On the 450-message simulated benchmark, the hybrid rule-LLM pipeline achieved $F1=0.906$ ($P=0.994$, $R=0.833$), substantially outperforming the rule-only baseline ($F1=0.567$, $R=0.397$). A fine-tuned BERT-base-Chinese established a supervised ceiling at $F1=0.965$ (5-fold CV). Performance varied by difficulty: easy 1.000, medium 1.000, hard 0.654—all 35 false negatives involved implicit causality or colloquial expressions. Four Chinese LLMs showed model-independent robustness ($F1: 0.892-0.950$). On 1,792 real clinical messages, the pipeline achieved $F1=0.905$ ($\Delta=-0.001$ vs. simulated) with zero false positives, and maintained 100% specificity on 3,897 negative control messages from a pharmacy quality-control group. These pilot results established feasibility but were limited by single-annotator benchmarks and the absence of temporal validation, motivating Phase 2.

3.2 Architecture Comparison

Table 1 presents the three pipeline architectures on both datasets.

Table 1: Pipeline architecture comparison on development (n=1,277) and locked test (n=746) sets. 95% bootstrap CI for F1 in brackets.

Dataset	Architecture	P	R	F1	95% CI
Dev	Rule-only	1.000	0.998	0.999	[0.998, 1.000]
	LLM-only	0.971	0.951	0.961	[0.952, 0.969]
	Hybrid	0.972	1.000	0.986	[0.981, 0.991]
Test	Rule-only	1.000	0.998	0.999	[0.998, 1.000]
	LLM-only	0.943	0.992	0.967	[0.957, 0.977]
	Hybrid	0.944	1.000	0.971	[0.961, 0.981]

256 The rule-only approach achieved F1=0.999 on both sets. **Importantly, this near-**
257 **perfect performance is setting-specific** to this pharmacovigilance-dedicated group
258 where standardized reporting conventions ensure high vocabulary coverage. On messages
259 with diverse expression patterns—including brand names, abbreviations (e.g., “MTX”),
260 and colloquial symptom descriptions—rule-only F1 dropped to 0.567 in Phase 1, because
261 vocabulary gaps cause missed detections. This 0.432 F1 gap between settings underscores
262 that rule-based approaches are not generalizable without extensive vocabulary engineer-
263 ing. The LLM-only configuration achieved F1=0.961 (95% CI: 0.952–0.969; Dev) and
264 0.967 (95% CI: 0.957–0.977; Test) without vocabulary engineering. The LLM-only vs.
265 hybrid difference was not significant on the test set (McNemar $\chi^2(1)=3.20$, $p=0.074$).
266 Given the LLM’s vocabulary independence, it was used as the primary configuration for
267 subsequent experiments.

268 3.3 Prompt Optimization and Temporal Validation

269 Three prompt strategies were compared on the development set (Table S3). The few-
270 shot balanced strategy achieved F1=0.961, far exceeding the strict negative (0.614) and
271 intermediate (0.830) strategies. Enhancement with confidence scoring and chemotherapy
272 guidance raised F1 to 0.977 (P=0.962, R=0.992; FN reduced from 52 to 9; 95% CI in
273 Table S3). On the locked test set, the enhanced prompt achieved F1=0.970 (95% CI:
274 0.960–0.978; P=0.944, R=0.997), confirming stable generalization ($\Delta=-0.007$).

275 3.4 Confidence Calibration and Two-Pass Ablation

276 Confidence scores were well-calibrated (ECE=0.030, Brier=0.039; Figure 2a). The distri-
 277 bution was strongly bimodal (Figure 2b): 85.6% of messages scored 8–10, 14.3% scored
 278 1–3, and only 0.1% fell in the uncertain range (4–7). Accuracy increased monotonically
 279 with confidence: 100% at score 1 (n=133) to 99.3% at score 10 (n=454).

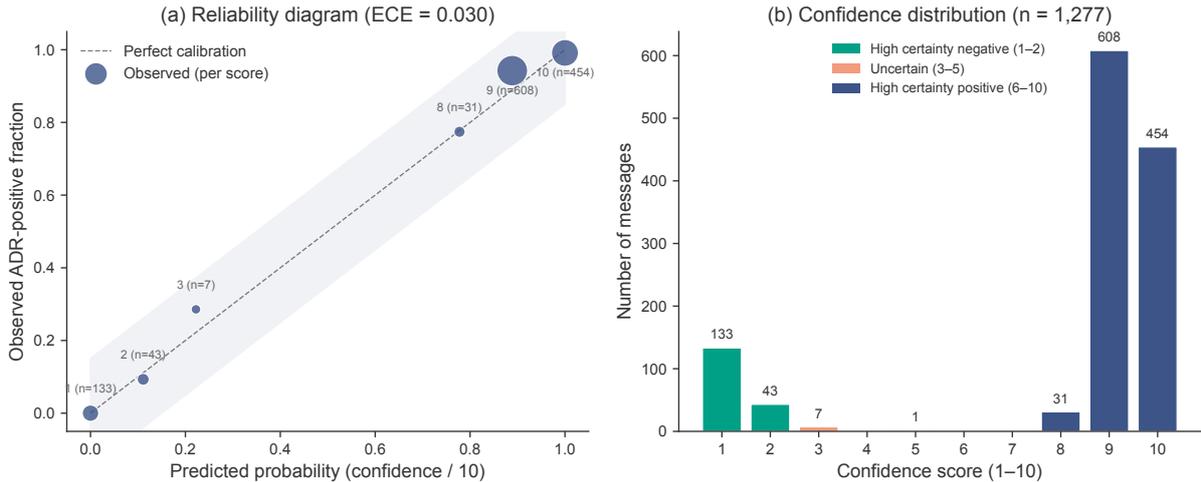


Figure 2: Confidence calibration on the development set (n=1,277). (a) Reliability diagram: each point represents one confidence score (1–10, mapped to predicted probability via $(s - 1)/9$; scores with $n < 5$ omitted), sized by sample count; the dashed line indicates perfect calibration and the shaded band marks ± 0.15 tolerance. (b) Confidence score distribution showing a bimodal pattern.

280 With optimized thresholds ($\theta_H=9$, $\theta_L=2$), 39 cases (3.1%) were routed to Pass 2.
 281 Two-pass F1=0.973, lower than single-pass F1=0.977 (McNemar $\chi^2(1)=1.57$, $p=0.211$),
 282 as re-examination introduced more errors than it resolved.

283 3.5 Multi-Turn Context Analysis

284 Single-turn classification achieved F1=0.967 on the development multi-turn subset (245
 285 conversations, 791 messages), significantly outperforming all context windows ($\chi^2(1)=12.90$ –
 286 19.53, all McNemar $p < 0.001$ after Bonferroni correction for four comparisons; Table S5).
 287 Context increased false positives (32 to 46–50) without recovering any of 9 single-turn
 288 false negatives. On the test subset, single-turn F1=0.968 (95% CI: 0.954–0.982).

289 **3.6 Cross-Model Comparison**

290 Four Chinese LLMs showed consistent performance on the development set (Table S11):
291 F1 ranged from 0.958 (Kimi K2.5) to 0.968 (GLM-5) with overlapping confidence intervals,
292 confirming that the pipeline’s effectiveness is not model-dependent. Five independent runs
293 at temperature=0.1 yielded CV=0.0005 with 99.2% unanimous agreement (Table S6).

294 **3.7 Specificity on Medical Negative Controls**

295 Overall specificity on 1,000 synthetic messages was 95.5% (Figure S6). Confounding
296 patterns were most challenging (72.5% specificity, 33/120 FP), reflecting the sensitivity–
297 specificity trade-off of the enhanced prompt. Seven of 11 categories achieved $\geq 98\%$ speci-
298 ficity.

299 Error analysis of 71 false positives on the real clinical data (development and test
300 sets combined) revealed three dominant patterns: chemotherapy-related expected ef-
301 fects (36.6%), drug–symptom co-occurrences without clear causal framing (31.0%), and
302 other borderline clinical narratives (29.6%). These patterns partially overlap with the
303 confounding-pattern category in synthetic controls, suggesting that drug–symptom co-
304 occurrence without causal framing is the primary challenge for the pipeline.

305 **3.8 Supervised Baseline**

306 A fine-tuned BERT-base-Chinese achieved F1=0.970 on the test set (Table S7), matching
307 the few-shot LLM pipeline (F1=0.970). The bootstrapped F1 difference was -0.001 (95%
308 CI: $[-0.014, 0.013]$), confirming non-inferiority within a ± 0.015 margin.

309 **3.9 Causality Assessment**

310 Automated Naranjo assessment showed no agreement with pharmacist consensus (Ta-
311 ble 2): $ICC(2,1)=-0.286$, $\kappa=-0.010$, $MAE=3.68$. Per-question analysis (Table S8) re-
312 vealed that 5 of 10 questions were answered “Unknown” in $\geq 99.5\%$ of cases. Restricting
313 to five IM-answerable questions did not improve agreement. The IMCT showed similarly

314 poor results ($\kappa=-0.135$, $ICC=-0.236$; Figure S4).

Table 2: Causality assessment: automated vs. pharmacist consensus (n=200).

Metric	Naranjo (10Q)	Naranjo (5Q)	IMCT (3Q)
ICC(2,1)	-0.286	-0.286	-0.236
Category κ	-0.010	-0.011	-0.135
MAE	3.68	3.67	N/A ^a

^aNot applicable; IMCT uses categorical triage (High/Medium/Low).

315 **3.10 Entity Extraction**

316 Entity extraction achieved lenient drug F1=0.780 (Dev) and 0.750 (Test), with lenient
317 symptom F1=0.764 and 0.738 respectively; 95% bootstrap CIs are reported in Table S12.
318 Evaluation was limited to ADR-positive messages with pharmacist entity annotations
319 (n=958); end-to-end performance on the full message stream—including true negatives
320 where no entity should be extracted—was not assessed. Strict match failures primarily
321 reflected systematic naming differences: the LLM translated abbreviations to full generic
322 names (e.g., “MTX”→“methotrexate”). Cross-model extraction was consistent (lenient
323 drug F1: 0.745–0.780; symptom F1: 0.735–0.764; Table S10).

324 Taken together, these results converge on three findings: (1) the pipeline achieves
325 near-expert ADR detection with well-calibrated confidence scores and model-independent
326 robustness; (2) ADR reports are self-contained within individual messages, making multi-
327 turn modeling unnecessary; and (3) automated causality assessment is structurally infea-
328 sible from IM data alone, regardless of instrument design. The Discussion examines the
329 mechanisms and implications of each finding.

330 **4 Discussion**

331 **4.1 Principal Findings**

332 This study validated a confidence-calibrated LLM pipeline for ADR detection from clinical
333 IM conversations. Three principal findings emerged.

334 First, the pipeline achieved near-expert classification with well-calibrated confidence.
335 On the locked test set, the LLM-only configuration reached $F1=0.970$ ($P=0.944$, $R=0.997$),
336 matching the supervised BERT baseline without task-specific training. Confidence scores
337 were well-calibrated ($ECE=0.030$, $Brier=0.039$) with a strongly bimodal distribution, pro-
338 viding a principled mechanism for deployment-time review prioritization: high-confidence
339 predictions (scores 8–10, comprising 85.6% of messages) can be accepted with minimal
340 oversight, while the rare uncertain cases can be routed for pharmacist review. Notably,
341 two-pass re-examination of uncertain cases did not improve classification accuracy (Mc-
342 Nemar $\chi^2(1)=1.57$, $p=0.211$), suggesting that the primary value of confidence scores lies
343 in triage rather than iterative refinement. Four independent Chinese LLMs all achieved
344 $F1 \geq 0.958$ with overlapping confidence intervals.

345 Second, ADR reports in this IM setting are self-contained within individual mes-
346 sages. Single-message classification significantly outperformed all multi-turn configura-
347 tions ($F1=0.967$ vs. $0.946-0.950$; $p < 0.001$), and the pipeline maintained 95.5% specificity
348 on non-ADR medical content.

349 Third, automated causality assessment showed no agreement with pharmacist con-
350 sensus across three instrument variants ($ICC \leq -0.236$), confirming that the barrier is
351 information asymmetry rather than instrument design.

352 4.2 Why Single-Turn Classification Outperforms Multi-Turn

353 The superiority of single-message classification contradicts the general expectation that
354 conversational context improves NLP tasks [8]. Three factors explain this. First, phar-
355 macists reporting suspected ADRs include drug name, symptom, and temporal framing
356 within a single message, making reports operationally self-contained. Second, preced-
357 ing messages often describe different patients, introducing irrelevant drug–symptom co-
358 occurrences that the LLM misinterprets as ADR signals—explaining the increased false
359 positives (32 to 46–50) with context. Third, the 9 single-turn false negatives involved
360 chemotherapy lab values and ultra-short messages requiring domain interpretation, not
361 contextual information.

362 This finding has broader design implications: IM-based pharmacovigilance systems on
363 dedicated reporting channels can adopt per-message classification, reducing complexity
364 and latency.

365 **4.3 Causality Assessment: A Structural Mismatch**

366 The Naranjo scale’s failure ($ICC=-0.286$) reflects a mismatch between its design—structured
367 case evaluation with complete medical records [21]—and the limited information in IM
368 messages. Five of 10 questions require data categorically absent from IM text, as con-
369 firmed by both LLM and pharmacists answering “Unknown” in $\geq 99.5\%$ of cases.

370 For the five remaining questions, poor agreement ($\kappa < 0.1$) reveals that pharmacists
371 draw on tacit clinical knowledge when interpreting brief messages, while the LLM ap-
372 plies this knowledge inconsistently. The IMCT, designed specifically for IM-answerable
373 dimensions, showed comparably poor results ($ICC=-0.236$). The consistency across all
374 three instruments demonstrates that the barrier is not instrument complexity but a fun-
375 damental information asymmetry between what pharmacists infer from clinical context
376 and what can be extracted from individual messages. IM-based pharmacovigilance should
377 therefore be scoped as detection and triage, with causality assessment deferred to formal
378 review.

379 **4.4 Prevalence and Deployment Considerations**

380 The high ADR prevalence in this pharmacovigilance-dedicated group (83.6%) warrants
381 discussion of lower-prevalence settings. Using the observed sensitivity (0.997) and speci-
382 ficity from synthetic medical negative controls (0.955), Bayesian projections (Figure 3)
383 yield: at 10% ADR prevalence, positive predictive value (PPV)=71% and negative pre-
384 dictive value (NPV) $>99.9\%$; at 5% prevalence, PPV=54% and NPV $>99.9\%$. These pro-
385 jections are based on specificity estimated from synthetic medical negative controls and
386 assume that this specificity transfers to authentic non-ADR clinical communication. Real-
387 world specificity may differ if authentic non-ADR messages exhibit patterns not repre-
388 sented in the synthetic corpus; validation on authentic non-ADR IM data is a priority

389 for future work. At 10% prevalence, approximately 7 of 10 flagged messages would be
390 true ADR reports while fewer than 1 in 3,000 unflagged messages would be missed. Con-
391 fidence scores provide an additional filtering mechanism—restricting to high-confidence
392 alerts (≥ 9) would increase PPV at the cost of modest recall reduction.

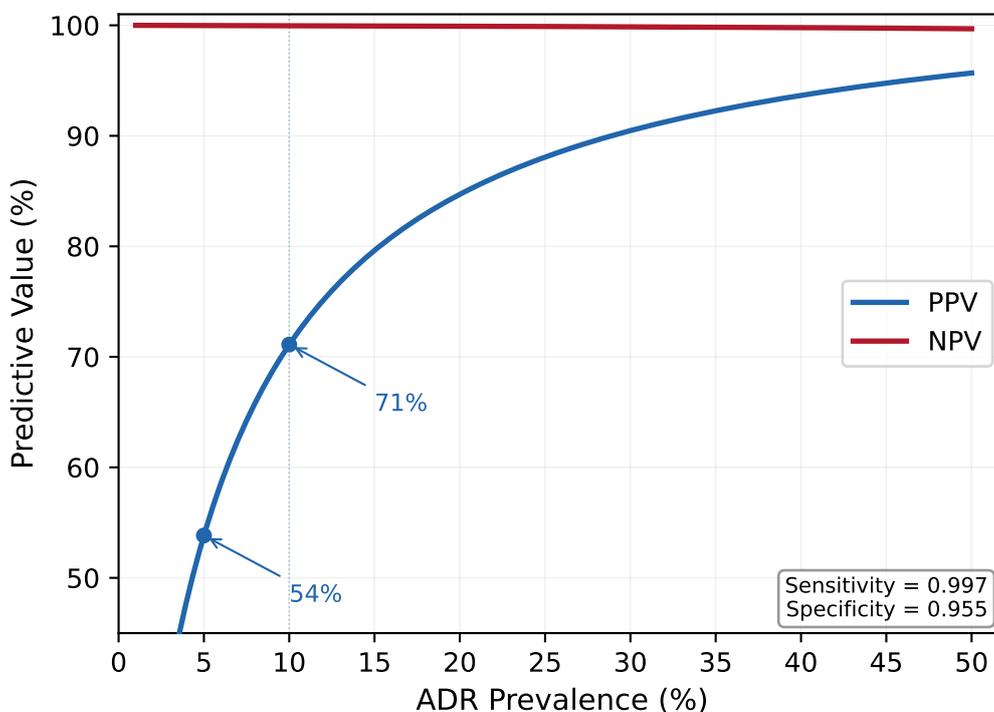


Figure 3: Prevalence-adjusted predictive values using observed sensitivity (0.997) and specificity (0.955). PPV increases steeply with prevalence, reaching 71% at 10% prevalence. NPV remains $>99.9\%$ across the displayed range.

393 In a practical deployment scenario, the pipeline would operate as a continuous mon-
394 itor on clinical IM groups, flagging suspected ADR messages for pharmacist review in a
395 dashboard ranked by confidence score. In a group generating approximately 50 messages
396 per day at 10% ADR prevalence, the system would flag an estimated 7 messages daily
397 (5 true ADRs, 2 false positives), representing a manageable review workload that could
398 augment existing spontaneous reporting workflows.

399 4.5 Comparison with Prior Work

400 Compared to the Phase 1 pilot, Phase 2 demonstrates four improvements: F1 increased
401 from 0.906 to 0.970 on temporally validated data; inference stability improved 10-fold

402 (CV: 0.0047→0.0005); the evaluation incorporated 5-annotator blind review with tem-
403 poral splitting; and entity extraction was validated against pharmacist annotations for
404 the first time. The few-shot LLM pipeline matched the supervised BERT ceiling (both
405 F1=0.970; bootstrapped $\Delta F1=-0.001$, 95% CI: [-0.014, 0.013]), demonstrating that
406 domain-specific prompt engineering with in-context examples eliminates the need for task-
407 specific fine-tuning.

408 In the broader NLP-based ADR detection literature, our results are consistent with
409 evidence that LLMs achieve near-expert performance on clinical NLP tasks [13, 14, 17].
410 However, this is the first demonstration on clinical IM text—a source distinct from elec-
411 tronic health records (EHRs), clinical notes, and social media in its brevity and real-time
412 conversational structure. While recent work has applied deep learning NLP to EHR
413 discharge summaries for automated ADR reporting [27], IM-based pharmacovigilance in-
414 volves unique challenges including ultra-short messages, absence of structured fields, and
415 informal language that differentiate it from EHR-based text mining.

416 4.6 Limitations

417 *Gold standard and adjudication.* The two-stage adjudication process changed 734 of 2,023
418 labels (36.3%) from ADR- to ADR+, substantially increasing prevalence from 47.4%
419 (majority vote) to 83.6%. While all adjudicated cases met the pre-specified objective
420 criteria (drug name, adverse symptom, and temporal association), this rate of override—
421 particularly 497 cases where 0/5 annotators initially voted positive—warrants scrutiny.
422 The high override rate suggests either that annotators systematically under-recognized
423 certain ADR patterns (e.g., dechallenge narratives, chemotherapy lab values) or that
424 adjudication criteria were overly inclusive. Future studies should consider adjudicator
425 blinding, formal inter-adjudicator agreement measurement, and bidirectional adjudica-
426 tion. The Naranjo gold standard used only three annotators for 200 cases; inter-annotator
427 ICC(2,1)=0.476 indicates moderate reliability for the gold standard itself, compounding
428 uncertainty in the LLM-vs-consensus ICC=-0.286 reported in Table 2.

429 *Generalizability and negative controls.* This single-center study at a tertiary obstetrics

430 and gynecology hospital limits external validity: the ADR spectrum is specialty-specific,
431 all evaluated models are Chinese-language LLMs, and the pharmacovigilance-dedicated
432 WeChat group has established reporting conventions that may not exist in general clinical
433 communication channels. Furthermore, the synthetic negative controls may not fully
434 capture authentic non-ADR clinical communication; validation on real non-ADR IM data
435 is a priority for future work. Multi-center, multi-specialty, and multi-language validation
436 is planned.

437 *Technical scope.* The near-perfect rule-only F1 (0.999) reflects this group’s standard-
438 ized conventions (F1=0.567 on diverse expressions in Phase 1) and should not be general-
439 ized. Entity extraction evaluation was limited to ADR-positive messages with annotations
440 (n=958), and enhanced recall on chemotherapy ADRs reduced confounding-pattern speci-
441 ficity to 72.5%. Retrieval-augmented generation with drug-specific knowledge bases may
442 improve discrimination.

443 *Model and temporal constraints.* All four evaluated models are Chinese-developed
444 LLMs accessed via cloud APIs; international models (e.g., GPT-4, Claude) were not
445 evaluated due to data residency requirements. API updates or model deprecation could
446 affect reproducibility, and the 3.5-year temporal split may not capture longer-term drift
447 in drug formularies or clinical practices.

448 *Interpretability.* The LLM provides natural language reasoning for each classifica-
449 tion, but this reasoning is generated alongside the decision rather than derived from a
450 transparent decision process. Feature-level attributions or attention-based explanations
451 that would allow clinicians to verify which message elements drove each classification are
452 not currently available. Future work should explore interpretability methods to increase
453 clinical trust in automated ADR detection.

454 5 Conclusion

455 A confidence-calibrated LLM pipeline achieves near-expert ADR detection from clinical
456 IM conversations (test F1=0.970), matching a supervised BERT baseline without task-
457 specific training, with well-calibrated confidence scores (ECE=0.030) and consistent per-

458 formance across four Chinese LLMs (F1=0.958–0.968). ADR reports are self-contained
459 within individual messages, eliminating the need for multi-turn modeling. Prevalence-
460 adjusted analysis supports deployment as a high-sensitivity screening tool (projected
461 PPV=71% at 10% prevalence, based on synthetic negative controls); these projections
462 and all findings require validation on authentic non-ADR IM data and across multiple
463 centers and specialties. Causality assessment remains fundamentally limited by the infor-
464 mation content of IM data regardless of instrument design ($ICC \leq -0.236$), indicating that
465 these systems should serve as detection and triage tools with causality assessment deferred
466 to formal review. Future work should prioritize multi-center validation and integration
467 with electronic health records for post-detection causality assessment.

468 **Acknowledgments**

469 The authors thank the five clinical pharmacists who contributed to the blind-review an-
470 notation and the three pharmacists who provided Naranjo causality consensus scoring.

471 **Author Contributions**

472 D.W. and Z.L. conceived the study, developed the system, conducted the experiments, and
473 wrote the manuscript. W.Y. contributed to data curation and the annotation study. K.Y.
474 contributed to software development. D.Y. contributed to data analysis. Y.Y. and S.J.
475 supervised the project and revised the manuscript. All authors reviewed and approved
476 the final manuscript.

477 **Funding**

478 This research received no specific grant from any funding agency in the public, commercial,
479 or not-for-profit sectors.

480 Conflicts of Interest

481 The authors declare no conflicts of interest.

482 Data Availability

483 The clinical messaging data contain protected health information and cannot be shared
484 publicly under China’s Personal Information Protection Law [26]. The pipeline source
485 code, evaluation scripts, rule vocabulary lists, prompt templates, synthetic negative con-
486 trol generation code, and all statistical analysis code are available at [https://github.](https://github.com/Patrick647/adr-detection-confidence-calibrated)
487 [com/Patrick647/adr-detection-confidence-calibrated](https://github.com/Patrick647/adr-detection-confidence-calibrated). A synthetic demonstration
488 dataset preserving the distributional characteristics of the original data is included to
489 enable independent verification. Researchers seeking access to the de-identified clinical
490 dataset may contact the corresponding author; access will be granted subject to institu-
491 tional data sharing agreements and ethics committee approval.

492 References

- 493 [1] World Health Organization. Safety of medicines: A guide to detecting and reporting
494 adverse drug reactions. WHO/EDM/QSM/2002.2, 2002.
- 495 [2] Ashish K. Jha, Itziar Larizgoitia, Carmen Audera-Lopez, Nittita Prasopa-Plaizier,
496 Hugh Waters, and David W. Bates. The global burden of unsafe medical care:
497 analytic modelling of observational studies. *BMJ Quality & Safety*, 22(10):809–815,
498 2013. doi: 10.1136/bmjqs-2012-001748.
- 499 [3] Ania Syrowatka, Wenyu Song, Mary G. Amato, Dinesh Chakraborty, Katharine Har-
500 ris, Thomas Hartvigsen, Gretchen P. Jackson, Clemens Scott Kruse, Kathleen M.
501 Mazor, Bill Mclean, et al. Key use cases for artificial intelligence to reduce the fre-
502 quency of adverse drug events: a scoping review. *The Lancet Digital Health*, 4(2):
503 e137–e148, 2022. doi: 10.1016/S2589-7500(21)00229-6.

- 504 [4] Rave Harpaz, Alison Perez, Herbert S. Chase, Raul Rabadan, George Hripcsak,
505 and Carol Friedman. Biclustering of adverse drug events in the FDA’s spontaneous
506 reporting system. *Clinical Pharmacology & Therapeutics*, 89(2):243–250, 2011. doi:
507 10.1038/clpt.2010.285.
- 508 [5] Lorna Hazell and Saad A. W. Shakir. Under-reporting of adverse drug reac-
509 tions: a systematic review. *Drug Safety*, 29(5):385–396, 2006. doi: 10.2165/
510 00002018-200629050-00003.
- 511 [6] Elena López-González, Maria Teresa Herdeiro, and Adolfo Figueiras. Determinants
512 of under-reporting of adverse drug reactions: a systematic review. *Drug Safety*, 32
513 (1):19–31, 2009. doi: 10.2165/00002018-200932010-00002.
- 514 [7] Philip Alexander Routledge. Improving the spontaneous reporting of suspected ad-
515 verse drug reactions: an overview of systematic reviews. *British Journal of Clinical*
516 *Pharmacology*, 89(8):2377–2385, 2023. doi: 10.1111/bcp.15791.
- 517 [8] Jiancheng Ye. Transforming and facilitating health care delivery through social net-
518 working platforms: evidences and implications from WeChat. *JAMIA Open*, 7(2):
519 ooae047, 2024. doi: 10.1093/jamiaopen/ooae047.
- 520 [9] Rachel M. Murphy, Joanna E. Klotowska, Nicolette F. de Keizer, Kitty J. Jager,
521 Jan Hendrik Leopold, Dave A. Dongelmans, Ameen Abu-Hanna, and Martijn C.
522 Schut. Adverse drug event detection using natural language processing: a scoping
523 review of supervised learning methods. *PLOS ONE*, 18(1):e0279842, 2023. doi:
524 10.1371/journal.pone.0279842.
- 525 [10] Su Golder, Dongfang Xu, Karen O’Connor, Yunwen Wang, Mahak Batra, and
526 Graciela Gonzalez Hernandez. Leveraging natural language processing and ma-
527 chine learning methods for adverse drug event detection in electronic health/medical
528 records: a scoping review. *Drug Safety*, 48(4):321–337, 2025. doi: 10.1007/
529 s40264-024-01505-6.

- 530 [11] Su Golder, Karen O'Connor, Yunwen Wang, Ari Klein, and Graciela Gonzalez Her-
531 nandez. The value of social media analysis for adverse events detection and pharma-
532 covigilance: scoping review. *JMIR Public Health and Surveillance*, 10:e59167, 2024.
533 doi: 10.2196/59167.
- 534 [12] Oladapo Oyeboade and Rita Orji. Identifying adverse drug reactions from patient re-
535 views on social media using natural language processing. *Health Informatics Journal*,
536 29(1):14604582221136712, 2023. doi: 10.1177/14604582221136712.
- 537 [13] Cheng Peng, Xi Yang, Aokun Chen, Kaleb E. Smith, Nima PourNejatian, Anthony B.
538 Costa, Cheryl Martin, Mona G. Flores, Ying Zhang, Tanja Magoc, Gloria Lipori,
539 Duane A. Mitchell, Naykky S. Ospina, Mustafa M. Ahmed, William R. Hogan, Eliz-
540 abeth A. Shenkman, Jiang Bian, Yonghui Wu, and Yi He. A study of generative
541 large language model for medical research and healthcare. *npj Digital Medicine*, 6:
542 210, 2023. doi: 10.1038/s41746-023-00958-w.
- 543 [14] Zhaoyue Sun, Gabriele Pergola, Byron C. Wallace, and Yulan He. Leveraging Chat-
544 GPT in pharmacovigilance event extraction: an empirical study. In *Proceedings*
545 *of the 18th Conference of the European Chapter of the Association for Computa-*
546 *tional Linguistics (EACL)*, volume 2, pages 344–357, St. Julian's, Malta, 2024. doi:
547 10.18653/v1/2024.eacl-short.30.
- 548 [15] Yiming Li, Jianfu Li, Jingcheng He, and Cui Tao. AE-GPT: using large language
549 models to extract adverse events from surveillance reports—a use case with influenza
550 vaccine adverse events. *PLOS ONE*, 19(3):e0300919, 2024. doi: 10.1371/journal.
551 pone.0300919.
- 552 [16] Fan Dong, Wenjing Guo, Jie Liu, Tucker A. Patterson, and Huixiao Hong. BERT-
553 based language model for accurate drug adverse event extraction from social media:
554 implementation, evaluation, and contributions to pharmacovigilance practices. *Fron-*
555 *tiers in Public Health*, 12:1392180, 2024. doi: 10.3389/fpubh.2024.1392180.
- 556 [17] Mazin Mamun Zitu, Daniel H. Owen, Max E. Hartranft, Jakob D. Schmit, Nicolas

- 557 Villegas, James C. Byrd, et al. Large language models for adverse drug events: A
558 clinical perspective. *Journal of Clinical Medicine*, 14(15):5490, 2025. doi: 10.3390/
559 jcm14155490.
- 560 [18] Joseph L. Fleiss. Measuring nominal scale agreement among many raters. *Psycho-*
561 *logical Bulletin*, 76(5):378–382, 1971. doi: 10.1037/h0031619.
- 562 [19] J. Richard Landis and Gary G. Koch. The measurement of observer agreement for
563 categorical data. *Biometrics*, 33(1):159–174, 1977. doi: 10.2307/2529310.
- 564 [20] An Yang, Baosong Yang, Binyuan Hui, Bo Zheng, Bowen Yu, Chang Zhou, Cheng-
565 peng Li, Chengyuan Li, Dayiheng Liu, Fei Huang, et al. Qwen2.5 technical report.
566 *arXiv preprint arXiv:2412.15115*, 2024.
- 567 [21] Claudio A. Naranjo, Usua Busto, Edward M. Sellers, Paul Sandor, Ivan Ruiz, E. A.
568 Roberts, E. Janecek, Carlos Domecq, and David J. Greenblatt. A method for estimat-
569 ing the probability of adverse drug reactions. *Clinical Pharmacology & Therapeutics*,
570 30(2):239–245, 1981. doi: 10.1038/clpt.1981.154.
- 571 [22] DeepSeek-AI. DeepSeek-V3 technical report. *arXiv preprint arXiv:2412.19437*, 2024.
- 572 [23] GLM Team, Aohan Zeng, Bin Xu, Bowen Wang, Chenhui Zhang, Da Yin, Dan
573 Roessler, Jie Xie, Jinghan Zhao, Kai Yu, et al. ChatGLM: a family of large language
574 models from GLM-130B to GLM-4 All Tools. *arXiv preprint arXiv:2406.12793*, 2024.
- 575 [24] Jacob Devlin, Ming-Wei Chang, Kenton Lee, and Kristina Toutanova. BERT: pre-
576 training of deep bidirectional transformers for language understanding. In *Proceed-*
577 *ings of the 2019 Conference of the North American Chapter of the Association for*
578 *Computational Linguistics: Human Language Technologies (NAACL-HLT)*, pages
579 4171–4186, 2019. doi: 10.18653/v1/N19-1423.
- 580 [25] Fabian Pedregosa, Gaël Varoquaux, Alexandre Gramfort, Vincent Michel, Bertrand
581 Thirion, Olivier Grisel, Mathieu Blondel, Peter Prettenhofer, Ron Weiss, Vincent
582 Dubourg, et al. Scikit-learn: machine learning in Python. *Journal of Machine*
583 *Learning Research*, 12:2825–2830, 2011.

- 584 [26] National People's Congress of China. Personal information protection law of the
585 people's republic of china, 2021. Effective November 1, 2021.
- 586 [27] Christopher McMaster, Julia Chan, David F. L. Liew, Elizabeth Su, Albert G. Frau-
587 man, Wendy W. Chapman, and Douglas E. V. Pires. Developing a deep learning nat-
588 ural language processing algorithm for automated reporting of adverse drug reactions.
589 *Journal of Biomedical Informatics*, 137:104265, 2023. doi: 10.1016/j.jbi.2022.104265.

Supplementary Material

S1. Prompt Template

The enhanced V3 prompt (few-shot balanced with confidence scoring) is summarized below. The complete prompt with all examples is available in the code repository.

System instruction (abbreviated): You are a clinical pharmacist reviewing instant messages from a hospital pharmacy group. Determine whether each message reports an adverse drug reaction (ADR). An ADR-positive message must contain: (1) an identifiable drug name (including brand names, abbreviations, or generic names); (2) a described adverse symptom or clinical sign; and (3) an explicit or implied temporal association. Chemotherapy-related laboratory abnormalities (e.g., bone marrow suppression after chemotherapy) with temporal framing should be classified as ADR-positive.

Output format: JSON with fields: `is_adr` (boolean), `confidence` (integer 1–10), `drug_name` (string), `symptoms` (string), `patient_id` (string), `reasoning` (string).

Examples: Four annotated examples are provided—two ADR-positive (explicit drug-symptom-temporal pattern; chemotherapy lab abnormality) and two ADR-negative (dosage consultation; disease symptom without drug involvement).

S2. Prompt Optimization

Table 3 presents the comparison of three prompt strategies on the development set (n=1,277). The few-shot balanced strategy was selected for all subsequent experiments. Enhancement with confidence scoring (1–10 integer scale) and domain-specific chemotherapy guidance further improved performance. Prompt versions V1–V2 were discarded during early development; V3–V5 represent the final candidate set.

S3. Medical Negative Control Categories

Table 4 provides descriptions of the 11 medical negative control categories with example messages.

Table 3: Prompt strategy comparison on development set (LLM-only, n=1,277).

Strategy	Description	P	R	F1	95% CI
Few-shot balanced	4 examples (2+, 2-)	0.971	0.951	0.961	[0.952, 0.969]
Intermediate	Relaxed criteria	—	—	0.830	—
Strict negative	High-specificity focus	—	—	0.614	—
Enhanced	+ confidence + chemo guidance	0.962	0.992	0.977	[0.966, 0.980]

Table 4: Medical negative control categories (n=1,000).

Category	n	Description
Confounding pattern	120	Drug + symptom, but expected pharmacological effect
Positive followup	100	Treatment success reports
Treatment planning	100	Future treatment discussions
Disease symptom	90	Symptoms without drug involvement
Dosage consultation	90	Drug dosing inquiries
Laboratory result	90	Lab values without drug-reaction framing
Medication order	90	Prescription and dispensing records
Drug information	80	General pharmacological queries
Drug interaction	80	Drug-drug interaction discussions
Medication inventory	80	Stock and supply management
Patient education	80	Patient counseling content

616 S4. Multi-Turn Context Analysis

617 Table 5 presents the multi-turn context window ablation on the development multi-turn
618 subset (245 conversations, 791 messages). Single-turn classification significantly outper-
619 formed all multi-turn configurations (McNemar $p < 0.001$ for all comparisons).

Table 5: Multi-turn context window ablation on development multi-turn subset (n=791). McNemar p -values compare each window to single-turn.

Window	P	R	F1	95% CI	FP	FN
0 (single)	0.949	0.985	0.967	[0.956, 0.976]	32	9
1	0.921	0.972	0.946	[0.932, 0.958]	50	17
3	0.926	0.975	0.950	[0.938, 0.962]	47	15
5	0.927	0.972	0.949	[0.935, 0.960]	46	17
Full	0.923	0.970	0.946	[0.933, 0.959]	49	18

620 Of 9 false negatives in the single-turn configuration, none were recovered by any multi-
621 turn context window (window sizes 1, 3, 5, and full conversation). The 9 missed messages
622 comprised: 5 chemotherapy-related lab values, 2 ultra-short messages, and 2 ambiguous

causality descriptions. In all cases, the classification error stemmed from the message content itself (domain-specific interpretation) rather than missing contextual information from preceding messages.

S5. Inference Stability

Table 6 presents the inference stability analysis across five independent runs on the development set (n=1,277) at temperature = 0.1.

Table 6: Inference stability across five independent runs (Qwen 3.5 Plus, temperature=0.1, n=1,277).

Run	TP	FP	FN	P	F1
1	1,001	29	53	0.972	0.961
2	1,003	30	51	0.971	0.961
3	1,001	30	53	0.971	0.960
4	1,002	29	52	0.972	0.961
5	1,004	30	50	0.971	0.962
F1 CV				0.0005	
Unanimous agreement				99.2% (1,267/1,277)	

Of 1,277 development set messages evaluated across five independent runs, 10 messages (0.8%) showed any classification disagreement. All 10 were borderline cases with short text length (median 18 characters) and ambiguous drug-symptom relationships. The maximum disagreement was 3/5 runs classifying as ADR-positive and 2/5 as ADR-negative, indicating that even unstable messages exhibited narrow decision boundaries.

S6. Supervised Baseline Comparison

Table 7 presents the fine-tuned BERT-base-Chinese supervised baseline alongside the few-shot LLM-only pipeline.

Table 7: Supervised BERT baseline vs. few-shot LLM pipeline on locked test set (n=746).

Model	Approach	P	R	F1	95% CI
BERT-base-Chinese	Fine-tuned (5-fold)	0.949	0.992	0.970	[0.961, 0.980]
Qwen 3.5 Plus	Few-shot LLM-only	0.944	0.997	0.970	[0.960, 0.978]

637 The fine-tuned BERT model achieved 5-fold cross-validation $F1 = 0.983$ ($SD = 0.007$)
638 on the development set, indicating strong in-distribution performance. On the locked
639 test set, both models achieved identical $F1 = 0.970$ with overlapping confidence intervals,
640 confirming that task-specific fine-tuning provides no additional benefit when domain-
641 specific prompt engineering is applied.

642 S7. Naranjo Per-Question Analysis

643 Table 8 presents the per-question agreement between automated LLM assessment and
644 pharmacist consensus for the 10 Naranjo questions.

Table 8: Naranjo per-question agreement between LLM and pharmacist consensus ($n=200$).
IM answerability reflects whether the question can typically be answered from IM text alone.

Q	Question (abbreviated)	Accuracy	κ	IM answerable
1	Prior ADR reports?	0.530	0.100	Medium
2	ADR after drug use?	0.270	-0.113	High
3	Improved after stopping?	0.555	0.006	High
4	Recurred on rechallenge?	0.995	0.000	Very low
5	Alternative causes?	0.080	-0.001	Low
6	Placebo response?	1.000	—	Very low
7	Toxic drug levels?	1.000	—	Very low
8	Dose-response?	1.000	—	Low
9	Prior similar reaction?	0.990	-0.005	Low
10	Objective evidence?	0.520	0.041	Medium

645 Questions 4, 6, 7, 8, and 9 were answered as “Unknown” by the LLM in $\geq 99.5\%$ of
646 cases, confirming their fundamental unanswerability from IM data. For Q6–Q8, both LLM
647 and pharmacists agreed on “Unknown,” yielding perfect agreement but no discriminative
648 value. The five IM-answerable questions (Q1, Q2, Q3, Q5, Q10) showed poor agreement
649 ($\kappa < 0.1$), reflecting the information asymmetry between what pharmacists infer from
650 clinical context and what can be extracted from IM messages.

651 S8. Naranjo Score Distributions

652 The LLM systematically underscored cases compared to pharmacist consensus, with the
653 modal category being “Possible” (63.5%) for the LLM vs. “Probable” (62.0%) for phar-

Table 9: Naranjo category distribution: LLM vs. pharmacist consensus (n=200).

Category	LLM (%)	Pharmacist (%)
Definite (≥ 9)	0.0	1.0
Probable (5–8)	6.5	62.0
Possible (1–4)	63.5	31.0
Doubtful (≤ 0)	30.0	6.0

654 macists. This reflects the LLM’s conservative interpretation of limited IM evidence, de-
655 faulting to “Unknown” for questions where pharmacists applied tacit clinical knowledge
656 to infer positive answers.

657 S9. IMCT Category Distribution

658 Figure 4 presents the distribution of IM Causality Triage (IMCT) categories comparing
659 automated LLM assessment with pharmacist gold standard.

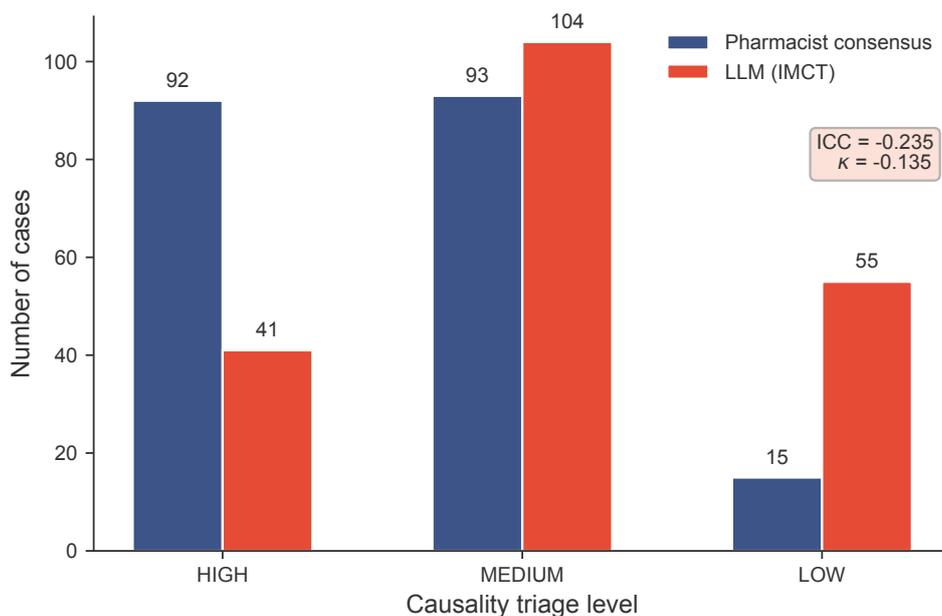


Figure 4: IMCT triage category distribution: automated LLM vs. pharmacist gold standard (n=200). The LLM systematically over-assigns Low (27.5% vs. 7.5%) and under-assigns High (20.5% vs. 46.0%), reflecting conservative interpretation of limited IM evidence.

660 **S10. Cross-Model Entity Extraction**

661 Table 10 and Figure 5 present entity extraction performance across four Chinese LLMs
 662 on the development set (n=615 ADR-positive messages).

Table 10: Cross-model entity extraction on development set (n=615 ADR-positive messages).
 Strict: normalized exact match; Lenient: token-overlap F1.

Model	Drug		Symptom	
	Strict	Lenient	Strict	Lenient
Qwen 3.5 Plus	0.761	0.780	0.803	0.764
GLM-5	0.699	0.776	0.836	0.761
Kimi K2.5	0.725	0.762	0.787	0.751
DeepSeek V3.2	0.698	0.745	0.772	0.735

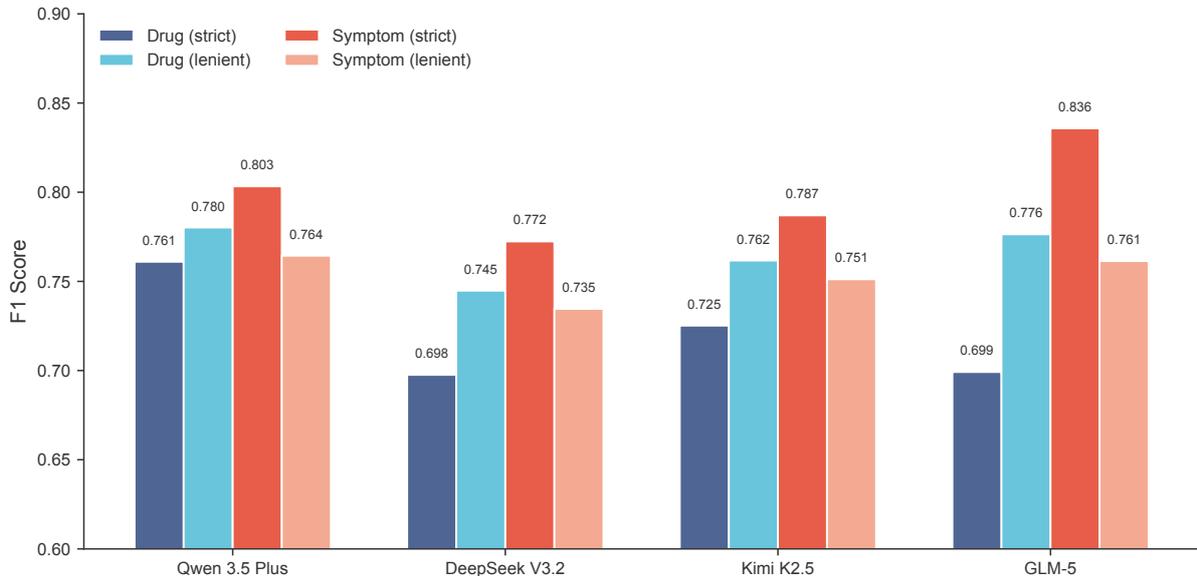


Figure 5: Cross-model entity extraction comparison on development set (n=615). All four models show consistent performance with lenient drug F1 in the range 0.745–0.780 and lenient symptom F1 in the range 0.735–0.764.

663 All four models achieved comparable extraction performance, with lenient drug F1
 664 ranging from 0.745 to 0.780 and lenient symptom F1 from 0.735 to 0.764. No single
 665 model dominated across both entity types: Qwen 3.5 Plus achieved the highest drug
 666 extraction while GLM-5 achieved the highest strict symptom match.

667 S11. Cross-Model ADR Classification

668 Table 11 presents the cross-model comparison for ADR classification on the development
669 set.

Table 11: Cross-model comparison on the development set (LLM-only, n=1,277).

Model	P	R	F1	95% CI	Spec
GLM-5	0.972	0.963	0.968	[0.960, 0.975]	0.870
DeepSeek V3.2	0.967	0.967	0.967	[0.959, 0.974]	0.843
Qwen 3.5 Plus	0.972	0.952	0.962	[0.952, 0.969]	0.870
Kimi K2.5	0.975	0.941	0.958	[0.948, 0.966]	0.883

670 S12. Entity Extraction

671 Table 12 presents entity extraction performance against pharmacist annotations.

Table 12: Entity extraction on ADR-positive messages (Dev n=615, Test n=343).

Entity	Split	n	Strict	Lenient	Empty
Drug	Dev	615	0.761	0.780	0.0%
	Test	343	0.706	0.750	0.3%
Symptom	Dev	615	0.803	0.764	0.2%
	Test	343	0.726	0.738	0.3%

672 Strict match failures primarily reflected systematic naming differences: the LLM trans-
673 lated abbreviations to full generic names (e.g., “MTX”→“methotrexate”) and provided
674 more specific symptom descriptions than the gold standard. For symptoms, strict match-
675 ing occasionally exceeded lenient matching (e.g., Dev: 0.803 vs. 0.764) because lenient
676 token-overlap scoring penalizes cases where the LLM provides a detailed multi-token de-
677 scription against a terse gold standard label, while strict matching after normalization
678 can still achieve an exact match.

679 S13. Specificity by Medical Negative Control Category

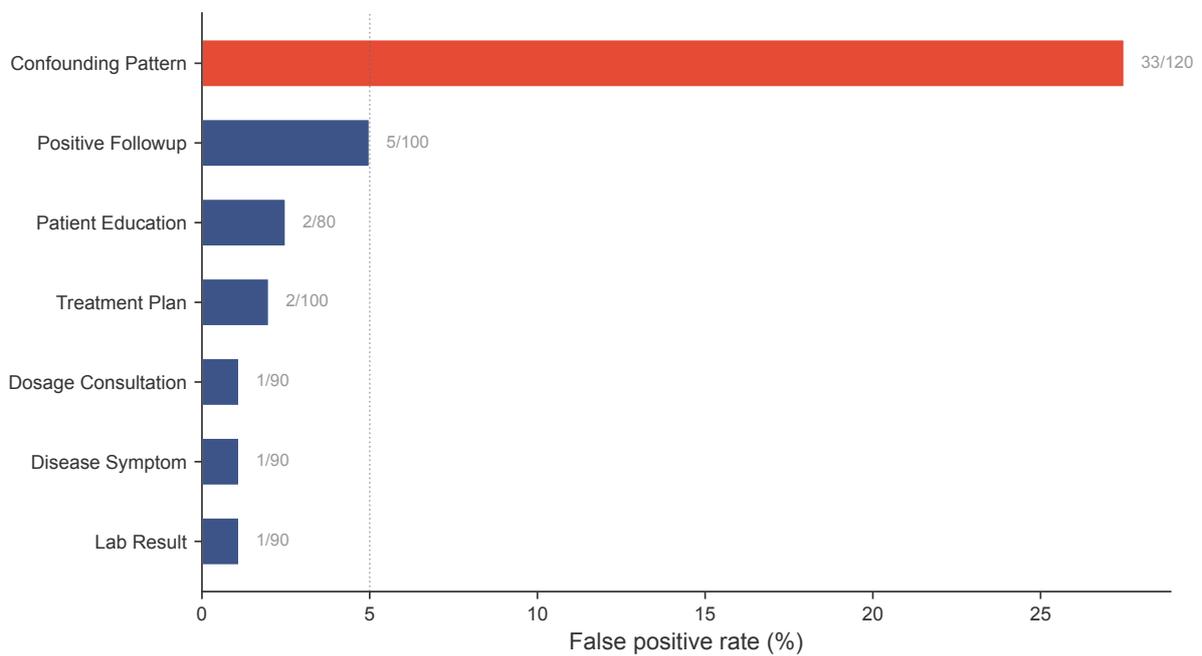


Figure 6: False positive rate by medical negative control category (n=1,000). Four categories with zero false positives (medication inventory, drug information, medication order, drug interaction) are omitted. Confounding patterns account for 73.3% of all false positives.